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Consumer Shopping and Spending Across Retail Formats

Abstract

We present an empirical study of household shopping and packaged goods spending across retail formats—grocery stores, mass merchandisers, and drug stores. Our study assesses competition between formats and explores how retailers' assortment, pricing, and promotional policies, as well as household demographics, affect shopping behavior. We find that consumer expenditures respond more to varying levels of assortment (in particular at grocery stores) and promotion than price. We also find that households that shop more at mass merchandisers also shop more in all other formats, suggesting that visits to mass merchandisers do not substitute for trips to the grocery store.

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Consumer Shopping and Spending across Retail Formats*

I. Introduction

Today, grocery stores operate in a competitive environment that includes other retail formats, in particular mass merchandisers. As a retail format, mass merchandisers grew rapidly throughout the 1980s and 1990s and currently generate nearly as much revenue as supermarkets.¹ The sale of groceries has traditionally been the venue of supermarket retailers like Kroger, Safeway, and Albertsons. However, mass merchandisers such as Wal-Mart, Target, and Kmart now offer thousands of packaged goods products that are also found in grocery stores. Additionally, other retail formats like drug stores such as Walgreens, CVS, and Eckerd also sell significant assortments of grocery

* We gratefully acknowledge Information Resources Inc., for providing the data used in this study. Contact the corresponding author, Edward J. Fox, at EFox@mail.cox.smu.edu.

1. Mass merchandisers (discount stores, warehouse clubs, and other mass merchants) reported 1998 sales of \$302.7 billion (computed from *Discount Store News* 1999) for both grocery and non-grocery items. Wal-Mart, the most prominent mass merchandiser, alone had sales in 1999 of \$137.6 billion. In comparison, supermarkets reported 1998 sales of \$346.1 billion (*Progressive Grocer Report of the Grocery Industry* 1999).

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We present an empirical study of household shopping and packaged goods spending across retail formats—grocery stores, mass merchandisers, and drug stores. Our study assesses competition between formats and explores how retailers' assortment, pricing, and promotional policies, as well as household demographics, affect shopping behavior. We find that consumer expenditures respond more to varying levels of assortment (in particular at grocery stores) and promotion than price. We also find that households that shop more at mass merchandisers also shop more in all other formats, suggesting that visits to mass merchandisers do not substitute for trips to the grocery store.

items. The overlap in their product offerings raises a fundamental question about competition across retail formats—do grocers and mass merchandisers compete for packaged goods sales? The grocery industry believes itself to be in direct competition with mass merchandisers. According to the *Progressive Grocer Report of the Grocery Industry* (1999), Wal-Mart represents a “grave” threat to grocery retailers. Moreover, this perceived threat has prompted consolidation and strategic changes among grocery retail firms.

Our study of shopping across retail formats is made possible by a panel data set collected by Information Resources Inc. (IRI), which is new to both industry and academic research. This data set is different from other shopping panels in that the panelists use scanners in their homes to record purchases at *all* retail formats and outlets. By contrast, other panels that are commonly analyzed in marketing use only purchases scanned in-store at participating supermarkets. The more complete household purchase records enable us to analyze shopping and spending across grocery stores, mass merchandisers, and drug stores with a rich set of predictors. Our analysis addresses how factors within and outside the retailers’ control affect store-level shopping decisions. We are also able to examine both intra- and interformat competition. Our goal is to measure and characterize this competition and also to consider its importance to retail managers. This article presents an exploratory analysis of consumer response across retail formats, which is intended to provide a foundation for future research in multiformat shopping behavior and retailer decision making in nongrocery formats.

There has been very little empirical research on shopping at mass merchandisers and other nongrocery formats, despite their growing importance, because of the lack of data on cross-format purchases. Store choice research has focused exclusively on grocery stores (Barnard and Hensher 1992; Bell, Ho, and Tang 1998; Bell and Lattin 1998; Ho, Tang, and Bell 1998). It is problematic to generalize from this work to other retail formats because grocery stores differ systematically from other formats in their marketing policies. For example, mass merchandisers offer lower prices, more product categories (e.g., groceries, clothing, garden, automotive products, etc.), smaller assortments within categories (i.e., fewer product variants), and fewer promotional discounts than grocers.

In this article, we estimate an econometric model to determine how marketing policies affect shopping and spending behaviors across retail formats. Our model incorporates consumer decisions about both “where to shop” and, conditional on shopping, “how much to spend.” The where to shop (patronage) decision is modeled as a binary choice for each store chain, which is correlated across chains. The conditional spending decision at each store chain is modeled as a continuous variable and also correlated across chains. In addition, we model differences between households because of known (i.e., demographic) and random factors with a Bayesian hierarchical specification. Thus, we develop and estimate a hierarchical, multivariate specification of the type-2 tobit model (using Amemiya’s [1985] topology) for this application. To our knowl-

edge this is the first such application of a multivariate type-2 tobit model, although univariate versions of the type-2 tobit and other multivariate tobit models have been published (Blundell and Smith 1994; Cornick, Cox, and Gould 1994). The rationale for specifying the type-2 tobit model is that it allows predictors to have different effects on the censoring decision and the continuous relationship, which in our case are where to shop and how much to spend.

The rest of this article is organized as follows: we describe the data in Section II and present our model in Section III. Section IV details our empirical analysis at both the store-chain and format levels and reports model fit and contribution of predictor variables. Section V considers the implications of our study for managers, and we conclude with a summary of our findings and opportunities for future research in Section VI.

II. Data Description

Before formally modeling shopping behavior across retail formats, we present a descriptive analysis of retailer marketing policies and consumer shopping behaviors across store chains in our data set. We use a panel data set that captures household-level shopping and spending across store chains and retail formats. This spending includes all items with uniform product codes (UPCs, or bar codes) that can be scanned, as well as non-scannable items like perishables (e.g., produce, meats, and bakery goods). We model the shopping behavior of 96 households at six different store chains representing three retail formats—grocery stores, drug stores, and mass merchandisers—over a 2-year period in a major U.S. market.² Within each format, the store chain(s) selected is (are) the largest and collectively account for the majority of the spending in that format.

Table 1 describes the marketing policies at each chain in our data set—average prices paid and regular (i.e., nonpromoted) shelf prices, promotional discounts, percentage of sales on promotion, within-category assortments, and travel times for shoppers.³ Price and assortment indices reflect the relative price of a basket of the most commonly purchased products (1,605 UPCs) and the relative number of products within the most commonly purchased categories (top 26 categories, or 10% of total), respectively. Formal definitions of these indices are given in Section III. Promotional discounts are the percentage off of the regular retail prices.

2. We eliminated households with suspect or incomplete information from our analysis. A household was omitted if either the majority of its purchases were made outside the chains included in this study or if the household did not record a purchase during any given month over the 2-year period of our analysis. This screen can potentially eliminate households that might have been on vacation for extended periods, but we believe that most of the households screened were not faithfully recording their purchases. The demographics of the remaining households were checked against the demographics of the zip codes in which they lived and found to be representative of these areas.

3. The entire data set, including those omitted as described in n. 2, is used to compute table 1.

TABLE 1 Descriptive Statistics for Marketing Policies across Store Chains and Retail Formats

	Price Index	Shelf Price Index	Promotional Discount (%)	Sales on Promotion (%)	Assortment Index	Travel Time (Min.)
Grocery 1	1.012	1.019	21.5	21.7	1.824	11.9
Grocery 2	1.023	1.016	18.6	14.7	1.913	10.1
Grocery average	1.017	1.017	20.1	18.2	1.868	11.0
Mass merchandiser 1	.950	.916	18.1	13.2	.647	14.1
Mass merchandiser 2	.921	.916	18.0	17.9	.602	15.2
Mass merchandiser 3	.918	.902	14.1	10.5	.513	17.6
Mass merchandiser average	.930	.912	16.7	13.9	.587	15.6
Drug store	.995	1.022	24.3	29.3	.408	9.6

Differences in marketing policies across retail formats are clearly evident. In particular, product assortment is much greater at grocery chains. In common product categories, grocers offer many more product alternatives than mass merchandisers, which in turn offer more product alternatives than drug stores. Roughly speaking, grocers offer more than three times the assortment of mass merchandisers and more than four times the assortment of drug stores. These extensive assortments are offered at a cost of either breadth of product variety (i.e., there are fewer, less-diverse product categories at grocery stores than mass merchandisers) or larger stores (i.e., grocery stores have more floor space than drug stores).

Mass merchandisers are the lowest-priced format, offering prices that average 7% and 9% less than drug and grocery stores, respectively, and regular shelf prices that average 11% and 10% less than drug and grocery stores, respectively. Promotional discounts are deepest at the drug store chain, followed in order by grocers and mass merchandisers. Similarly, the highest percentage of promotional sales are made at drug stores, followed in order by grocers and mass merchandisers. Although regular shelf prices at drug stores are higher than other formats, drug store patrons still appear price sensitive as demonstrated by the higher percentage of sales on promotion. Clearly, drug store patrons make use of the deep promotional discounts provided.

Moreover, one should not conclude that price-sensitive consumers will patronize mass merchandisers, the lowest-priced format, inordinately. Even though the average prices at grocery stores may be higher, their more extensive assortments and deeper discounts may provide more fertile ground for search than mass merchandisers, making them attractive to price-sensitive shoppers. Likewise, the even deeper discounts at the drug store chain may also attract price-sensitive shoppers. We also note that grocers' larger assortments may increase shopping time, so time-constrained households may be willing to trade off the convenience of drug stores for the higher prices. Finally, the larger number of drug stores results in the lowest average travel times from

TABLE 2 Descriptive Statistics for Shopping Behaviors across Store Chains and Retail Formats

	Interval between Trips (Days)	Total Spending per Trip (\$)	Patronage— Households Shopping per Month (%)
Grocery 1	10.1	85.70	58.2
Grocery 2	7.1	88.65	83.6
Grocery store average	8.6	87.18	70.9
Mass merchandiser 1	29.1	78.71	28.4
Mass merchandiser 2	13.6	80.52	36.9
Mass merchandiser 3	18.3	84.09	33.2
Mass merchandiser average	15.9	81.11	32.9
Drug store	19.7	36.36	46.0

shoppers' households, followed in order by grocery stores and mass merchandisers.

Table 2 describes shopping behaviors at each chain in our data set: the interval between shopping trips, spending per trip, and patronage at that chain. We define patronage as the percentage of households shopping at a store chain in a given month. Approximately one-third of households frequent a given mass merchandiser each month, while almost half visit the drug store, and more than 70% visit the grocer. The time interval between shopping trips follows a slightly different pattern. It is shortest for grocery chains, with an average shopping interval of just over 1 week (8.6 days), while visits to mass merchandisers or drug stores occur every 2–3 weeks. Households spend substantially more on trips to grocery stores (\$87.18) and mass merchandisers (\$81.11) than on trips to drug stores (\$36.36). Recall that all purchases, not only packaged goods, are reflected in these spending statistics. In summary, customers shop more frequently and spend more at grocery stores than at other formats.

The relationships of marketing policies and shopping behaviors to retail format are quite strong. Most of the variation is between formats, while variation within formats is relatively small. First, we consider marketing policies. An analysis of variation reveals that 99%, 99%, 93%, 84%, and 76% of the variance for assortment, regular shelf price, actual price, travel time, and promotional discount (given in table 1), respectively, are accounted for between formats, as opposed to within format.⁴ The strong relationships between marketing policies and retail format suggest that consumer response to marketing variables may differ by format, which we explore in subsequent analyses. Another analysis of variance shows that 99%, 83%, and 58% of the variation in shopping behaviors—spending per trip, patronage, and the interval between trips, respectively (given in table 2)—is explained by format. Overall

4. All ANOVAs reported in this section are one-way analyses that use format to predict chain-level averages over the period of our data.

then, we find that both marketing policies and shopping behaviors are format specific. This implies that retail format is a good segmentation criterion for examining retailers.

III. The Model

In order to relate retailer marketing policies to shopper patronage and spending decisions, we develop relative measures of price, promotion, and assortments that incorporate the many products included in shoppers' market baskets. These measures are constructed by averaging price, promotional status, and product availability over all packaged goods products, weighted by the household's long-term category consumption rate as measured by total consumption during a 2-year period. This assumes that store-level (as opposed to product-level) shopping decisions depend on the relative price of the shopper's entire market basket at different stores (e.g., Bell, Ho, and Tang 1998; Bell, Bucklin, and Sismiero 2000). We then use these price indices to predict monthly expenditures at each retail chain.

Perhaps the most controversial aspect of our modeling approach is temporal aggregation. The alternative would be to model shopping on a trip-by-trip basis or perhaps weekly. While a more disaggregate analysis may be generally preferred, we believe that it is not practical using our data set. We offer several reasons why temporal aggregation is desirable for this application.

First, a trip-level model would require the complete set of prices and promotions offered by each retailer. Unfortunately, this information is not available, nor is it even possible to get such information from IRI, since key mass merchandisers forbid the distribution of this information by IRI on confidentiality grounds. Hence, the sparseness of marketing policy data for mass merchandisers is not conducive to trip-level modeling.⁵ We are forced to piece together the casual data, as summarized in table 1, from panelists' observed purchases (587,279 individual packaged goods purchases across 261 categories in the six store chains over 2 years [see n. 3]). Fortunately, we have determined that price variation is very robust to temporal aggregation from the weekly to the monthly level (see app. A).⁶

Second, a trip-level model would require information about consumption,

5. Disaggregate data for price, promotion, and assortment variables are particularly sparse for drug stores and mass merchandisers. For example, of the 2,000 products that they purchase most frequently, our panelists purchase only 52 and 22 products weekly at the average mass merchandiser and drug store chain, respectively. Creating price and promotional measures that reflect the many products therefore requires pooling observations over multiple weeks.

6. To understand why price variability is stable over temporal aggregation, consider that the majority of items remain at regular shelf price throughout the month, and unadvertised promotions persist for multiple weeks to nearly a month (4 weeks) in this market. The prices that do change weekly, feature advertised items, are fixed in number by space limitations in the advertising circular. Thus, the assortment of feature advertised items change weekly, but their number is relatively constant. Weekly advertised discount depth is also fairly constant over time (mean = 23.5%, SD = 3.8%). As a result, price variation because of advertised promotions is not affected by temporal aggregation.

household inventory levels, and nongrocery items—information that we lack. While others have attempted to infer inventory as a parameter of a purchasing model (Bell, Ho, and Tang 1998), the lack of full causal information in our model would likely make these estimates unstable. This problem is exacerbated in our data set, since mass merchandising trips may be triggered by nongrocery purchases, for which we lack item-level data. For example, a consumer may buy clothing products (which are not tracked by UPC) and, while at the store, decide to purchase several grocery items that are either low in the household's inventory or a good value. Hence, in our data set we would have difficulty explaining how consumers plan their mass merchandiser visits. Furthermore, we would expect that purchase (and consumption) will be more consistent at an aggregate level, since it is possible that week-to-week purchases could be quite volatile because of heavy workloads or vacations that might lead to more eating out and less grocery purchases. Therefore, without adequate information about consumption, inventory, and nongrocery items, an individual trip-level model would be difficult to estimate.

Third, our purpose is to predict consumer behavior at an aggregate level. When the true underlying generating model is unknown and nonlinear, temporal aggregation can result in a more linear model (Man 2004). It is quite likely that individual trip-level shopping behavior will depend nonlinearly on inventory levels, price expectations, and planned consumption. Hence, temporal aggregation can help simplify the model structure. Diagnostics of our model's residuals do not show significant autocorrelation; hence, our choice of monthly aggregation would seem to simplify the model.

There are disadvantages to estimating our model at an aggregate rather than a disaggregate level. Foremost, the interpretation of our variables is more difficult, since many tactical decisions are made at a weekly level. Aggregation can result in the loss of information when specifying and estimating the model; the predictions of aggregate models are not as good as aggregating predictions from a disaggregate model when the true generating process is known. Moreover, knowledge about structural properties of utility and demand could be utilized better at a disaggregate level. While we grant that a disaggregate model of cross-format retail shopping would be desirable, given the nascent state of research about retail formats, we believe that it is better to propose a simpler model that makes weaker assumptions than a more complex one whose assumptions cannot be readily validated. Therefore, we believe that the disadvantages are outweighed by the benefits of aggregation. We hope that the insights into the aggregate process will help future researchers in understanding the aggregate properties that disaggregate models of cross-shopping behavior must satisfy.

A. *Model Specification*

The household's spending decision at the store chain of interest is modeled as a regression model with the log of the household's monthly expenditures

at that store chain as the continuous dependent variable. Our data consist of 13,824 observations (purchases of 96 households at six chains over 24 months). More than half of our expenditure observations, 7,227 out of 13,824, are zeros. Following Tobin (1958), we treat our continuous spending variable as censored; that is, this regression is conditioned on a binary probit model for whether or not the chain was visited.

The variable of interest in our model, y_{hit} , is the expenditures made by household h (indexed $h = 1, \dots, H$; $H = 96$) at chain i ($i = 1, \dots, S$; $S = 6$) during month t ($t = 1, \dots, T$; $T = 24$). Expenditures are observed only when an indicator variable for household h 's patronage at chain i at time t , z_{hit} , takes the value of one. The observational equation for expenditures, y_{hit} , is

$$y_{hit} = \begin{cases} y_{hit}^* & \text{if } z_{hit} = 1 \\ 0 & \text{otherwise} \end{cases}, \quad (1)$$

where the model for the logarithm of the latent variable, y_{hit}^* , is

$$\ln(y_{hit}^*) = \alpha_{hi} + \mathbf{x}'_{hit}\boldsymbol{\beta}_i + d_{hi}\gamma_i + \mathbf{s}'_i\boldsymbol{\lambda}_i + \varepsilon_{hit}. \quad (2)$$

The incidence of store patronage, z_{hit} , is a binary variable, and we use a probit model to describe its behavior:

$$z_{hit} = \begin{cases} 1 & \text{if } z_{hit}^* \geq 0 \\ 0 & \text{otherwise} \end{cases}. \quad (3)$$

The latent variable, z_{hit}^* , is modeled through a linear model:

$$z_{hit}^* = \iota_{hi} + \mathbf{x}'_{hit}\boldsymbol{\theta}_i + d_{hi}\chi_i + \mathbf{s}'_i\boldsymbol{\kappa}_i + \mathbf{u}_{hit}. \quad (4)$$

The predictors \mathbf{x}_{hit} , d_{hi} , and \mathbf{s}_i used in equations (2) and (4) are the same. However, these predictors may influence a shopper's patronage decision differently from her spending decision, so we allow for different coefficients in these two equations. Consider promotions as an example. Some retailers advertise everyday low prices yet in fact offer substantial discounts in order to generate store traffic (i.e., shopper patronage). However, the increased patronage may come from opportunistic shoppers who will spend less than more loyal shoppers. The vector of marketing policy variables, \mathbf{x}_{hit} , that applies to household h during month t at chain i comprises price, promotional intensity, and product assortment. The predictor d_{hi} is the travel time for household h to the nearest store of chain i , and \mathbf{s}_i is an 11×1 vector of indicator variables for the 12 months of the year, capturing seasonal effects.

As the subscripts of the intercept terms in equations (2) and (4) show, every household has an individual intercept coefficient for each store chain. In this

way, the model incorporates individual differences in preference for retailers.⁷ These preferences are modeled as a function of known and unknown (random) factors. This is accomplished by using a hierarchical specification for the intercept terms. Like Ainslie and Rossi's (1998) brand-choice model, we specify household-level preferences to be systematically affected by household characteristics, that is, demographics. For expenditure model intercepts,

$$\alpha_{hi} = \mathbf{w}_h' \boldsymbol{\delta}_i + \xi_{hi}, \quad (5)$$

and for patronage model intercepts,

$$\iota_{hi} = \mathbf{w}_h' \boldsymbol{\psi}_i + \tau_{hi} \quad (6)$$

The vector \mathbf{w}_h is a vector of household demographics—family size, income, home ownership, working woman, education, and the presence of children who are under age 6. A detailed discussion of these predictors is presented in the next subsection. Parameter vectors $\boldsymbol{\delta}_i$ and $\boldsymbol{\psi}_i$ relate household h 's demographics to its intrinsic preference for store chain i in the conditional spending and patronage models, respectively.

Our model incorporates both the binary choice of store patronage (i.e., will the household shop at Wal-Mart?) and the continuous decision of how much to spend at that store chain, given patronage. Note that the second decision is not spending per trip, but total spending during a given month at the chain. Thus, it may include multiple trips. Equations (1)–(4) define a type-2 tobit specification. Equations (5) and (6) define a hierarchical specification of individual preferences. We link the hierarchical type-2 tobit models for each chain in a multivariate framework by allowing errors to be correlated. The vector of household residuals from the log expenditure models, $\boldsymbol{\varepsilon}_{ht} = (\varepsilon_{h1t}, \varepsilon_{h2t}, \dots, \varepsilon_{hSt})'$, is assumed to follow a multivariate normal distribution: $\boldsymbol{\varepsilon}_{ht} \sim \text{MVN}(0, \boldsymbol{\Sigma})$. The vector of household residuals from the patronage models is $\mathbf{u}_{ht} = (u_{h1t}, u_{h2t}, \dots, u_{hSt})'$ and is also assumed to follow a multivariate normal distribution: $\mathbf{u}_{ht} \sim \text{MVN}(0, \boldsymbol{\Lambda})$. Because equations (2) and (4) use the same predictors, we must assume that $\boldsymbol{\varepsilon}_{ht}$ and \mathbf{u}_{ht} are independent in order to identify the model.

The multivariate error distributions allow information from one chain to influence the conditional predictions of another (see app. B). We expect prediction errors for conditional spending models to be correlated across store chains, because excess expenditures in one store should result in less spending at other stores. If the total budget for groceries is fixed during a month, then any increase in expenditures in one chain should lead to a decrease in purchases at other chains, which in our model would be captured through a negative correlation. We also considered including cross-effects of the marketing mix

7. Jain, Vilcassim, and Chintagunta (1994) find that most household-level heterogeneity in brand choice results from differences in intrinsic brand preferences. By allowing for household-specific preferences for store chains, we control for unmodeled individual differences, including average basket size (Bell and Lattin 1998). We thank an anonymous reviewer for raising this point.

variables directly in equations (2) and (4), but, because of the multicollinearity of this specification, we allow cross-effects to enter only through the error covariance matrix. One can think of our model specification as a cross-effects model in reduced form. Our specification of error correlations for equations (2) and (4) improves the efficiency of parameter estimation as in seemingly unrelated regression (Zellner 1962).

We also allow for relationships in a household's intrinsic preferences for different store chains by specifying error correlations for the hierarchy—equations (5) and (6). The vector of residuals for equation (5), which models household-level preferences for store spending, $\xi_{ht} = (\xi_{h1t} \ \xi_{h2t} \ \cdots \ \xi_{hSt})'$, is assumed to follow a multivariate normal distribution: $\xi_{ht} \sim \text{MVN}(0, \mathbf{V}_\alpha)$. Similarly, the residual vector for equation (6), $\tau_{ht} = (\tau_{h1t} \ \tau_{h2t} \ \cdots \ \tau_{hSt})'$, which models household-level preferences for store patronage, is assumed to follow a multivariate normal distribution: $\tau_{ht} \sim \text{MVN}(0, \mathbf{V}_t)$.

Only recently have Markov Chain Monte Carlo (MCMC) methods for estimating the posterior distributions of the parameters in high dimensional models become available (Gelfand and Smith [1990]; Casella and George [1992]; also see Chib [1993], for application to the tobit specification). A complete discussion of our estimation procedure is available as a technical report from the authors.

B. Predictors of Shopping Behavior

To predict expenditures across households and stores, we have defined three sets of variables. First, we consider retailer marketing policies: pricing, promotion, and product assortment. These variables are firmly under the retailer's control and, with the exception of assortment, are easily manipulated in the short-term. Second, we measure the costs incurred by the shopper traveling to and from the store. This variable is also affected by the retailer, but only by its long-run store location decisions and market penetration. Third, demographic characteristics of the shopper's household (e.g., family size, income, and home ownership), which are not affected by the retailer, are included in the hierarchy. All predictors are own effects. As noted, we are unable to incorporate cross-effects (i.e., marketing policies at one retailer are specified as predictors of shopping behavior at another) directly into the models because of multicollinearity.

1. *Retailer marketing policies.* Previous research on store choice and store sales has shown the importance of retailer prices and promotions on shopping behavior (Arnold, Ma, and Tigert 1978; Arnold and Tigert 1982; Arnold, Oum, and Tigert 1983; Walters and Rinne 1986; Kumar and Leone 1988; Walters and MacKenzie 1988; Walters 1991; Barnard and Hensher 1992; Bell, Ho, and Tang 1998; Bell and Lattin 1998). Our PRICE variable is a price index weighted by the household's long-term consumption, which captures variation in expected prices across chains and individual households. This construction is similar to the construction of Dillon and Gupta's (1996)

household-specific category price variable, which weighted brand prices by each household's long-term brand consumption. In our application, a price index is used to capture variation in the price of the household's average market basket across both stores and time. The composition of a household's average market basket is based on weighted-average category consumption over a 2-year period and so is unlikely to be affected by short-term price or promotional variation (Ainslie and Rossi 1998).⁸

Formally, the price index, PRICE_{hit} , for household h at store i in month t is the weighted average of category-level price indices at store i for category c at time t , which is denoted as p_{ict} (recall that these indices are based on observed panelist purchases). Category price indices are weighted by household h 's long-term consumption of products in each category c , denoted q_{hc} . Thus, we have:

$$\text{PRICE}_{hit} = \frac{1}{C} \sum_{c=1}^C \left(\frac{p_{ict}}{\bar{p}_{ct}} \frac{q_{hc}}{\sum_{c=1}^C q_{hc}} \right), \quad (7)$$

where \bar{p}_{ct} is the average price of products in category c at time t across store chains.

As noted, different market baskets are used for each household, incorporating purchases across all stores and time periods. Thus, the market basket is specific to the household, but not to the month or store chain. This allows PRICE and other expenditure-weighted variables to reflect differences in long-term consumption between households. To illustrate, consider one household that has an infant child and buys diapers and baby food. A second household has no children and never purchases these categories. The prices of diapers and baby food will be reflected in the price indices applied to the first household at every store chain that offers them. Prices of these categories will be weighted in proportion to the first household's total expenditures on diapers and baby food. The prices of diapers and baby food will not be reflected in the price indices applied to the second household. In this way, the model captures variation in PRICE and other marketing policies across households based on differences in market baskets between households.

Retailer promotions are also well known to affect shopping behavior (see Blattberg, Briesch, and Fox [1995] for a review). The PROMO variable summarizes the retailer's promotional policies. It is operationalized as the proportion of all purchases at a store chain that are made on promoted items, again weighted according to the household's average market basket. In this

8. This argument is advanced and then tested by Ainslie and Rossi (1998), who use household-level shopping behavior variables to predict brand choice. They note that "these variables are computed as long-run averages of shopping behavior in which bursts of promotional activity will be averaged out" (p. 97). They report results with and without shopping behavior variables and find no evidence of endogeneity.

way, both the frequency and depth of retailer promotions are incorporated into a single variable.

Another marketing policy that has been shown to affect shopping behavior and patronage patterns is product assortment (Reilly 1931; Huff 1964; Brown 1989).⁹ The ASSORT variable is an indexed measure of the number of products within each category, weighted by the household's average market basket. It therefore reflects the diversity of the product offering within category, not across categories. It is constructed like the other indexed variable, PRICE. Because assortment varies little over time, variation in ASSORT is virtually all cross-sectional.

2. Travel time. Virtually all models of retail competition (Hotelling 1929; Reilly 1931; Huff 1964; Hubbard 1978; Brown 1989) and shopping behavior (Barnard and Hensher 1992; Arentze, Borgers, and Timmermans 1993; Bell, Ho, and Tang 1998; Dellaert et al. 1998) specify store patronage as a function of the distance from the store to the shopper's home. Our model includes a measure of distance in the form of travel time, TRAVTIME, which is operationalized as the time in minutes it takes to travel from the household to the nearest store of a given chain.¹⁰ The underlying assumption is that the shopper travels from home to the closest store of the selected chain, then returns home. In reality, shoppers may reduce their travel time by linking shopping trips together or combining store visits with other required travel. "Trip chaining," as this practice is called (Thill and Thomas 1987), results in shoppers requiring less than the measured travel time to make a store visit and possibly shopping more than expected at distant stores. We expect measurement error because of trip chaining to bias the estimated effect of travel time downward.

The models also include household-specific intercepts for each store chain. The intercept reflects the intrinsic preference for that particular store chain, including unobserved factors specific to that retailer, which affect shopping behavior. These factors include the retailer's general positioning (e.g., high service, friendly), operational policies, and overall excellence in execution. They also include the variety, or breadth, of product categories offered. For example, mass merchandisers sell consumer durables and clothing not available in grocery or drug stores, while grocery stores offer perishable products that cannot be purchased in drug stores or mass merchandisers. Preference for this variety (which does not vary over time) is reflected in the intercept.

9. Following Levy and Weitz (2003), we consider assortment as the depth of the product offering in a category (e.g., the number of products offered per category), as opposed to variety, which is the breadth of the product offering (i.e., the number and diversity of categories offered). Variety is captured, along with other unobserved chain-specific variables, in the chain intercept term.

10. In our data set, distance is measured from the centroid of the zip + 4 in which the household is located to the street address of the closest store in the chain using a closest road algorithm. This measurement of distance is superior to that used in previous research because the panelist locations are considerably more precise and because the road distance more effectively captures the shopper's expected travel. The TRAVTIME variable is based on road distance and is adjusted for expected driving speed and traffic.

TABLE 3 Descriptive Statistics for Household Demographics

	Mean	SD
Family size (no.)	2.98	1.39
Income (× \$1000)	52.1	25.7
Working woman (%)	65.6	47.5
College educated (%)	15.6	36.3
Home owner (%)	86.5	34.2
Children under age 6 (%)	16.7	37.3

In addition, the intercept captures preference for mean levels of the retailer's marketing policies. For example, mass merchandisers consistently offer lower basket prices compared with grocery and drug stores, independent of short-term variation. Because the intercept terms reflect the household's response to these many factors, we do not attempt to interpret their coefficients, *per se*.

3. *Household characteristics.* Equations (5) and (6) specify household-level intercept coefficients that have a deterministic component based on known demographic characteristics. Previous research using cross-sectional (Blattberg et al. 1978; Hoch et al. 1995) and panel data (Ainslie and Rossi 1998) suggests that demographics can influence price sensitivity. We include the following demographic variables in our model: (1) income (INCOME), measured in thousands of dollars; (2) family size (FAMSIZE), which is the number of household members; (3) home ownership (OWNHOME), which is an indicator (1 = yes); (4) education (COLLEGE), which is an indicator (1 = yes); (5) working adult female (FEMWORK), which is an indicator (1 = yes); and (6) the presence of a young child (ages 0–6) in the home (YOUNGKID), which is an indicator (1 = children present). By including demographic variables in our hierarchical specification, we identify systematic sources of heterogeneity in patronage and spending across households. Descriptive statistics for the demographics of households in our data set are shown in table 3.

IV. Empirical Results

We estimate equations (1)–(6) for each chain simultaneously using a MCMC approach. Equations (2) and (4) result in two sets of parameter estimates. The first panel of table 4 shows coefficients for the probit component of the model, that is, whether the household will patronize a given store chain. The signs of coefficients relate positively to the patronage probability. Alternately, the derivative of the patronage probability with respect to the *j*th variable can be computed as $\phi(x'\theta)\theta_j$, where $x'\theta$ is the predicted value, and θ_j is the *j*th element of the parameter vector θ . Hence, the derivative is not constant and is influenced by the behavior of $\phi(x'\theta)$. The second panel of table 4 shows coefficients for the continuous component of the model, that is, how much the household

TABLE 4 Marketing Mix and Travel Time Parameter Estimates

	Drug Store	Grocery 1	Grocery 2	Mass Merchandiser 1	Mass Merchandiser 2	Mass Merchandiser 3
Patronage—where to shop:						
INTERCEPT	-.201*	.826***	3.284***	-1.159***	-1.011***	-.555***
	(.104)	(.169)	(.467)	(.122)	(.114)	(.130)
PRICE	.012	-.048	.153	.210***	-.087**	.070
	(.036)	(.120)	(.105)	(.063)	(.036)	(.061)
PROMO	-.025	-.144	.510*	.167**	.239***	.424***
	(.039)	(.183)	(.203)	(.057)	(.056)	(.076)
ASSORT	.055	.526***	.550***	-.017	.535***	-.001
	(.039)	(.113)	(.138)	(.051)	(.062)	(.079)
TRAVTIME	-.111**	-1.085***	-.375*	-.201***	.011	-.400***
	(.044)	(.160)	(.173)	(.068)	(.052)	(.086)
Expenditure—how much to spend:						
INTERCEPT	3.596***	4.646***	5.189***	4.628***	4.404***	4.512***
	(.101)	(.076)	(.059)	(.188)	(.133)	(.131)
PRICE	-.018	.008	-.019	-.066	.044	.056
	(.036)	(.042)	(.021)	(.074)	(.047)	(.070)
PROMO	-.191***	-.102*	.217***	.082	.283***	-.165*
	(.039)	(.057)	(.028)	(.077)	(.072)	(.085)
ASSORT	-.045	.280***	.191***	-.117*	.316***	.286**
	(.041)	(.066)	(.023)	(.060)	(.043)	(.089)
TRAVTIME	-.065*	-.304***	-.044**	-.037	-.011	.021
	(.041)	(.089)	(.017)	(.096)	(.045)	(.089)

NOTE.—Estimates are posterior means, with standard errors given in parentheses below each estimate.

* Statistically significant at .05.

** Statistically significant at .01.

*** Statistically significant at .001.

will spend, given that they patronize that retailer. The derivative of the log expenditure component with respect to the j th variable is β_j .

A. Parameter Estimates

1. *Marketing variables and travel.* Of the variables that we consider, PRICE is the weakest predictor of shopping and spending behavior. This may seem surprising, given that survey research finds price to have a substantial negative effect on store patronage (Arnold et al. 1978; Arnold and Tigert 1982; Arnold, Oum, and Tigert 1983). However, the six store chains in our data set all have very consistent pricing profiles through time, resulting in little variation in retailers' comparative basket prices. In fact, the average coefficient of variation for the basket price variable, PRICE, across those store chains is only 0.022. The price dispersion commonly observed in brand-choice modeling attenuates over the many products that make up the market basket. Appendix A demonstrates that aggregation across products reduces price variation far more than temporal aggregation. As Bell, Ho, and Tang (1998, p. 354, n. 3) argue, promotions effectively cancel one another out over the many items in the market basket. Because variation in PRICE within retailers is small, the effects of mean PRICE levels are captured through the intercept coefficients, along with other retailer-specific factors.

Prior research also helps us understand why price variation has little impact on store patronage and spending. Hoch, Dreze, and Purk (1994) found that consumers are inelastic to price changes for grocery purchases, which is consistent with our findings. Kalyanaram and Little (1994) demonstrate that consumers are not affected by small differences in price, provided that prices are close to their expectations. Moreover, unadvertised promotions, which comprise the majority of discounts, cannot be observed by the shopper until she visits the store. As a result, patronage decisions do not incorporate variation resulting from unadvertised discounts. Consumers also encounter difficulties in applying price information in basket shopping decisions. Alba et al. (1994) show that cognitive limitations result in consumers making erroneous comparisons of basket prices across retailers, based primarily on frequency cues. In sum, while visit-to-visit price variation strongly influences brand-level purchase decisions, the stability of basket prices, coupled with consumers' difficulty in learning basket prices for use in shopping decisions, explains why PRICE variation has little effect on consumer patronage and conditional spending.

The variable PROMO has a positive effect on patronage, with four of six store models positively signed and significant. In particular, promotions in categories of interest positively affect patronage at mass merchandisers. A priori, we expect that PROMO would have a positive effect on store expenditures, that is, consumers spend more as the depth and number of promotions in categories of interest increase. The effect of promotions on expenditures is mixed, however, with two significant positive and three significant negative

parameter estimates. Note that expenditures respond negatively to promotions in categories of interest at the retailers with highest promotional intensity, the drug store chain and grocery 1 (see table 1). One possible explanation is that retailers with high promotional intensity draw a disproportionate number of “cherry pickers” (customers who shop opportunistically across multiple stores during a single period) who buy less at each store. An alternative explanation is that, while store promotions may succeed in attracting additional shoppers to the store, conditional spending falls because demand is inelastic.

The variable ASSORT also has a positive effect on both patronage and spending. Across the two decisions, seven of the eight significant coefficients are positively signed. We note that assortment has the greatest effect at grocery stores, with positive and highly significant parameter estimates for patronage and spending at both stores. It may appear surprising that assortment has such a large effect on shopping behavior at grocery stores, given the grocery industry’s recent focus on reducing retail assortments (Information Resources and Willard Bishop Consulting 1993; *Food and Beverage Marketing* 1994; Merrefield 1995) and recent academic research, which finds that reduced online assortments lead to sales increases (Boatwright and Nunes 2001). However, evidence presented in these citations focuses on category sales and does not address the effect of assortment on patronage or store choice. In fact, Boatwright and Nunes (2001, p. 60) note “a significant decrease in the category purchase probability despite the increase in overall sales,” which they acknowledge is likely because of customer attrition.

The variable TRAVTIME has a substantial negative effect on store patronage, with five of the six coefficients negative and significant; TRAVTIME also has a negative effect on spending, though the effect is limited to drug and grocery stores. At mass merchandisers, differences in travel time across households that patronize the format do not affect their expenditures. This may be caused by the higher expenditures per trip at these low-priced stores (Fox, Metters, and Semple 2003) offsetting the reduction in trips because of longer travel times.

2. Demographics. The parameter estimates for equations (5) and (6) are shown in table 5. Coefficients of these hierarchical equations capture the systematic effects of consumer demographics on the intercept terms, ι_{hi} and α_{hi} , for patronage and spending models, respectively. The first panel of table 5 shows coefficients of the probit (i.e., patronage) models. The second panel shows the coefficients for the continuous regression (i.e., conditional spending) models. In general, we find relatively weak relationships between intercept coefficients and household demographics. Only 21 of 84 total coefficients (two consumer decisions \times six stores \times seven variables) are significant. This is not surprising, given limited prior success in relating demographics to category-level consumer decisions (Bucklin and Gupta 1992; Chintagunta and Gupta 1994; Rossi, McCulloch, and Allenby 1996). The variable FAMSIZE has the largest effect on store preferences, with six of the 12 intercept co-

efficients significant. A priori, we would expect larger households to spend more, because they have more members. We find that this is true for mass merchandisers, as evidenced by all positive coefficients (most significant) for patronage and spending. This suggests that larger households are more likely to patronize and spend more at mass merchandisers, which offer lower basket prices but fewer promotions.

The variables FEMWORK, COLLEGE, INCOME, and YOUNGKID are expected to increase shoppers' opportunity cost of time (Blattberg et al. 1978; Hoch et al. 1995). We expect that shoppers with higher opportunity costs of time will shop at fewer chains. The patronage model coefficients offer some support for this expectation for FEMWORK and COLLEGE. All of the FEMWORK patronage coefficients are negative, though none are significant. Most of the COLLEGE coefficients in the patronage models are also negative, as are the two significant coefficients. Of the patronage coefficients for INCOME and YOUNGKID, only one is significant and no patterns emerge. Overall, the variables that reflect opportunity cost of time appear to have limited effects on patronage and no differential influence across formats. Turning to the expenditure models in the second panel, we find that all FEMWORK spending coefficients are positive, and three are significant. We conclude that households with working women spend more at each retailer they patronize, though they patronize fewer retailers. This suggests that households with working women may be more loyal. Spending coefficients for COLLEGE, INCOME, and YOUNGKID offer few insights. The only significant coefficients for these variables are positive, but there are few. Interestingly, grocery 1 has positive and significant YOUNGKID coefficients for both the patronage and expenditure models. We conjecture that this retailer has a unique offering for young children, such as a "Baby Club," or successfully differentiates its stores by effectively merchandising categories such as diapers and baby food.

Home ownership (OWNHOME) is commonly interpreted as a proxy for storage space (Blattberg et al. 1978; Hoch et al. 1995). Shoppers with more space are able to "stock up," taking advantage of promotions, so that they can visit more chains in search of deals. Interestingly, we find that the drug store chain is significantly less likely to be visited by home owners. This is inconsistent with the above rationale, because the drug chain offers the deepest promotional discounts (see table 1), which home owners could exploit because of their storage space. By contrast, drug stores are a convenience format, not typically associated with "stocking up," and they carry limited product variety and assortment. It appears that the capability of home owners to stockpile packaged goods reduces their need for the convenience that drug stores offer—they do not stock up at drug stores despite the extensive promotions. Perhaps home owners use their storage space to exploit the one-stop-shopping benefits of broader-line grocery stores (Messinger and Narasimhan 1997) and mass merchandisers.

TABLE 5 Demographic Parameter Estimates from Hierarchy

	Drug Store	Grocery 1	Grocery 2	Mass Merchandiser 1	Mass Merchandiser 2	Mass Merchandiser 3
Patronage—where to shop:						
INTERCEPT	1.395* (.746)	-.365 (1.353)	4.217 (3.707)	-.979* (.684)	.355 (2.095)	-2.035 (1.435)
FAMSIZE	-.082 (.227)	-1.576*** (.488)	1.084 (1.163)	.524** (.234)	2.243** (.801)	.485 (.473)
INCOME	-.138 (.214)	.104 (.452)	.208 (1.145)	-.219 (.187)	.277 (.655)	-.183 (.433)
OWNHOME	-1.355* (.706)	.930 (1.250)	2.871 (3.109)	1.109* (.620)	-.022 (1.916)	1.712 (1.238)
FEMWORK	-.348 (.449)	-.649 (.910)	-1.296 (2.323)	-.255 (.385)	-1.827 (1.383)	-.608 (.868)
COLLEGE	-1.061* (.591)	-1.380 (1.112)	1.600 (3.287)	-.758 (.594)	-3.487* (1.943)	.296 (1.172)
YOUNGKID	.430 (.593)	3.092** (1.148)	-.749 (3.073)	-.477 (.527)	-1.980 (1.860)	1.038 (1.166)

Expenditure—how much to spend:							
INTERCEPT	.024	-.685*	.013	-.453*	.173	-.466	
	(.263)	(.370)	(.354)	(.227)	(.273)	(.330)	
FAMSIZE	-.005	-.180	.359***	.136*	.330***	.053	
	(.083)	(.123)	(.113)	(.070)	(.088)	(.106)	
INCOME	-.030	-.040	.063	-.015	.209*	-.147	
	(.086)	(.105)	(.111)	(.070)	(.091)	(.112)	
OWNHOME	-.311	.431	-.124	.256	-.166	.156	
	(.234)	(.339)	(.319)	(.222)	(.256)	(.298)	
FEMWORK	.365*	.365	.245	.259*	.086	.374*	
	(.167)	(.238)	(.230)	(.125)	(.179)	(.198)	
COLLEGE	-.102	-.109	-.124	.542**	-.392	.298	
	(.226)	(.291)	(.296)	(.180)	(.248)	(.259)	
YOUNGKID	.125	.563*	-.262	-.099	-.167	.216	
	(.228)	(.323)	(.315)	(.171)	(.221)	(.256)	

NOTE.—Estimates are posterior means, with standard errors given in parentheses below each estimate.

* Statistically significant at .05.

** Statistically significant at .01.

*** Statistically significant at .001.

B. Model Specification Testing and Contribution of Predictor Variables

To assess our model specification, we estimate nested models, which represent less general variants of our hierarchical multivariate type-2 tobit. Fit statistics for these models are shown in table 6. We estimate a system of independent type-2 tobit models (model A in table 6) as our baseline model. This baseline specification incorporates neither the individual differences modeled by the hierarchy nor the estimation efficiencies because of multivariate error structures. We assess the contribution to fit of multivariate error structures by estimating a multivariate type-2 tobit specification (model B). This specification offers an improvement in log likelihood of 892 ($u^2 = 0.054$), compared with the baseline. The incremental fit gained from specifying multivariate error structures suggests that there are meaningful unmodeled relationships among retailers in patronage and expenditures.

We then assess the incremental contribution of individual differences by estimating both a multivariate type-2 tobit with random intercepts (model C) and the full hierarchical multivariate type-2 tobit (model D), which allows for both systematic and random differences between households. Allowing random intercepts improves log likelihood by 1,205 over the model with fixed intercept (model B), with a u^2 of 0.126. The full hierarchical model offers a log-likelihood improvement of only 163 compared with the random intercepts specification, and a u^2 of 0.136. Clearly, individual differences attributable to household demographics offer only a fraction of the explanation available from unmodeled factors. Taken together, the large improvement in fit of models C and D over fixed-effects models suggests that individual differences offer substantial explanation of shopping behavior.

We also use nested models to assess the relative contributions of marketing variables and travel time. We do this by restricting travel time parameters to zero in model E, then restricting parameters for marketing variables to be zero in model F. By comparing the fit of these models with the full model, we find that the marketing variables make a substantially greater contribution to fit than travel time. The full model (D) is 223 log-likelihood points better than model E, but only 11 log-likelihood points better than model F.

Minimizing Akaike's Information Criterion favors selection of the full model, while minimizing Bayesian Information Criterion, which imposes a more severe penalty for additional parameters, favors a specification without travel time (model F). We include travel time because our objective is to make inferences about the effects of such factors on shopping behavior, even at the expense of parsimony.

C. Format-Level Empirical Results

The capability to relate individual-level decisions and parameters to retailer revenues is an important benefit of our model specification. Even though our model is estimated at the chain level, we can use it to predict revenue response at the format level. To do so we make two transformations. First, we focus

TABLE 6 **Nested Model Tests**

Model	Predictors	Individual Differences	Error Structure	Parameters	Log Likelihood	u^2	BIC	AIC
A	Marketing, travel, season	None	Independent	96	-16,601	.000	33,599	33,394
B	Marketing, travel, season	None	Multivariate	96	-15,709	.054	31,815	31,609
C	Marketing, travel, season	Random intercepts	Multivariate	102	-14,504	.126	29,430	29,212
D	Marketing, travel, season	Hierarchy with demographics	Multivariate	138	-14,341	.136	29,254	28,959
E	Travel, season	Hierarchy with demographics	Multivariate	120	-14,564	.123	29,625	29,368
F	Marketing, season	Hierarchy with demographics	Multivariate	132	-14,352	.135	29,250	28,967

NOTE.—BIC = Bayesian Information Criterion; AIC = Akaike's Information Criterion.

on expected revenues by (1) combining household patronage and conditional spending decisions into expected expenditures, and then (2) summing expected expenditures over households. Next, we sum expected revenues over store chains in each format in order to gain a better understanding of how the format as a whole behaves.

The parameters discussed in the previous subsection can be difficult to interpret individually since they separate the effect on shopper patronage from the effect on conditional spending. To summarize the two effects, we compute revenue elasticities. A revenue elasticity is defined as the average percentage change in the expected revenues of a format (based on expenditures of all households in the sample) in response to a 1% increase in the j th predictor variable (v_j). For example, our price elasticity of revenue for mass merchandisers measures the percentage change in revenues that would result from a 1% increase in prices, across all categories at each mass merchandiser. Note that the effect of such an across-the-board price increase would be independent of the household-level consumption weights.

Expected revenues incorporate expectations of both the probability of patronage and conditional expenditures, summed across all households h . First, we define the expected revenues at chain i :

$$E(R_{it}) = \sum_H E(y_{hit}^*) \Pr(z_{hit} = 1), \quad (8)$$

where

$$E(y_{hit}^*) = \alpha_i + \mathbf{x}_{hit} \boldsymbol{\beta}_i + d_{hi} \gamma_i + \mathbf{s}'_i \boldsymbol{\lambda}_i; \quad (9)$$

$$\Pr(z_{hit} = 1) = \Phi(\iota_{hi} + \mathbf{x}'_{hit} \boldsymbol{\theta}_i + d_{hi} \chi_i + \mathbf{s}'_i \boldsymbol{\kappa}_i). \quad (10)$$

We now aggregate the chain revenues to the format level:

$$E(R_{ft}) = \sum_{i \in F} E(R_{it}), \quad (11)$$

where F defines the set of retailers in format f . Finally, format-level elasticities for variable v are defined as:

$$\eta_{vft} = \frac{\partial E(R_{ft})}{\partial v_{ft}} \frac{v_{ft}}{E(R_{ft})}. \quad (12)$$

Our estimates of format-level revenue elasticities are given in table 7. The elasticities are estimated directly from the draws of the MCMC sample. Because elasticities are pure numbers, their magnitudes are comparable across formats.

1. *Price.* Revenue elasticities of price show that mass merchandisers would likely gain revenue by raising prices, while drug and grocery stores would lose revenues. However, none of the elasticities are significantly different than zero—not surprising given the minimal price variation and insignificant parameter estimates (see Sec. III.A). It is instructive to point out the

TABLE 7 Format-Level Elasticity Estimates

	Drug Store	Grocery Store	Mass Merchandiser
PRICE	-.372 (.945)	-1.050 (1.531)	1.299 (1.274)
PROMO	-1.323*** (.282)	.633*** (.153)	.729*** (.273)
ASSORT	-.109 (.212)	4.972*** (.554)	.897* (.413)
TRAVTIME	-.243* (.163)	-.308*** (.089)	-.106 (.156)

NOTE.—Standard errors are shown in parentheses below the estimates.

* Statistically significant at .05.

** Statistically significant at .01.

*** Statistically significant at .001.

difference between revenue price elasticities and the more commonly studied quantity price elasticities. Revenue price elasticity is equal to one plus the corresponding quantity price elasticity.¹¹ Therefore, the fact that revenue price elasticities are not significantly different from zero indicates that the corresponding quantity price elasticities are not significantly different from unity. This is consistent with the findings of Hoch et al. (1994), who find that grocers tend to face an inelastic aggregate demand curve.

2. *Promotion.* Average PROMO elasticities are highly significant in all formats, though the effect differs by format. The most promotionally oriented format, drug stores, would gain substantial revenues by promoting less ($\eta = -1.323$). We infer that the deep promotions generate neither the additional trips nor the additional spending per trip needed to offset the revenue lost on promotional discounts. This is consistent with the format's convenience positioning. By contrast, grocery stores ($\eta = 0.633$) and mass merchandisers ($\eta = 0.729$) would realize additional revenues by increasing promotions. We do not suggest that these formats would profit by increasing promotions, but consumer spending would increase. Note that revenue elasticities of promotion are inversely related to average promotional intensity. The highly promotional drug store chain (average discount of 23.4%) has a negative elasticity, while the less-promotional formats, grocery stores (average discount of 20.1%) and mass merchandisers (average discount of 16.7%), are positive. We will develop this observation further in the next subsection.

3. *Assortment.* Assortment elasticities are highest for grocery stores, 4.972, compared with 0.897 and -0.109 for mass and drug chains, respectively. The high sensitivity of grocers' revenues to product assortment levels is consistent with their actual assortment levels, which are far greater than

11. Let p, q, r represent price, quantity, and revenue, respectively, and the relationship between these variables is $r = pq$. Differentiating both sides with respect to price and reexpressing in percentage change terms, we find that $(\partial r/\partial p)(p/r) = 1 + (\partial q/\partial p)(p/q)$. In other words, the price elasticity of revenue equals one plus the price elasticity of quantity. Obviously, this holds true only for a single product, although we can think of our product in this case as a composite of all consumer packaged goods.

other formats (see table 1). These elasticities imply that if assortments were uniformly higher in grocery stores, revenues would also be substantially greater. However, given the floor-space constraints that grocers face, increasing assortments could prove challenging (and costly). Moreover, much of the grocery industry's recent focus has been on reducing retail assortments to gain cost advantages and operational efficiencies. Our results do not support this approach, suggesting that smaller assortments are associated with substantially lower revenues among grocery stores.

4. *Travel time.* Travel time has a consistent negative effect across formats, implying that as travel time increases to stores of each format, expenditures at that format decrease. Average elasticities for all formats are negative (two are significant), and the magnitudes of travel elasticities range from -0.308 to -0.106 . Grocers and drug stores have similar mean elasticity estimates, which are both significant. The lower revenue elasticity estimate for mass merchandisers suggests that the format is less sensitive to travel times. Again, this is likely because consumers stockpile goods when making trips to more distant, but lower-priced mass merchandisers. This stockpiling offsets the reduced probability of visiting more distant stores.

In summary, we find significant differences in revenue response across formats to marketing and travel variables. Grocery store revenues are highly sensitive to increases in category assortments. Their revenues also increase with increases in promotional intensity and decreases in shoppers' travel times. Mass merchandiser revenues are more sensitive than grocers to increases in promotion, but less sensitive than grocers to increases in assortment. Mass merchandisers suffer less than other formats from longer travel times, which generally result from lower market penetration. Finally, revenues of drug stores, the most promotional format, would benefit from decreasing promotional intensity. Drug stores suffer from increasing consumer travel times almost as much as grocers, but would not benefit at all from increasing category assortments.

D. Patterns of Cross-Format and Intraformat Shopping

The hierarchical specification of our multivariate type-2 tobit model generates household-level intercept coefficients, which represent consumers' intrinsic preferences for the store chains and formats being modeled. Understanding how preferences for stores are related within and across formats helps us to understand patterns of retail competition. We compute correlations in preferences among stores using draws from the MCMC sample and, thus, develop empirical distributions of store-preference correlations. The 6×6 store chain correlation matrices for patronage and spending decisions are reported in table 8.

The correlations suggest similarities and differences in patterns of competition for the two decisions. We begin with preference correlations for the patronage decision in the upper panel of table 8. The highest magnitude correlation is between grocery chains ($\rho = -0.633$). This large negative cor-

TABLE 8 Preference Correlations across Chains

	Drug Store	Grocery 1	Grocery 2	Mass Merchandiser 1	Mass Merchandiser 2	Mass Merchandiser 3
Patronage—where to shop:						
Drug store	1	.273** (.051)	-.063 (.066)	.210*** (.050)	.003 (.049)	.064 (.048)
Grocery 1		1	-.633*** (.054)	-.058 (.073)	-.082 (.052)	.076 (.059)
Grocery 2			1	-.036 (.071)	.141* (.069)	-.012 (.084)
Mass merchandiser 1				1	.309*** (.047)	.137** (.049)
Mass merchandiser 2					1	.238*** (.052)
Mass merchandiser 3						1
Expenditure—how much to spend:						
Drug store	1	.217** (.074)	-.019 (.061)	.563*** (.077)	-.093 (.098)	-.032 (.130)
Grocery 1		1	-.363*** (.047)	.180 (.118)	-.011 (.068)	.045 (.102)
Grocery 2			1	.246*** (.104)	.518*** (.055)	-.127* (.073)
Mass merchandiser 1				1	.339* (.122)	-.185 (.121)
Mass merchandiser 2					1	.013 (.101)
Mass merchandiser 3						1

NOTE.—Standard errors are shown in parentheses below the estimates.

* Statistically significant at .05.

** Statistically significant at .01.

*** Statistically significant at .001.

relation suggests that the more consumers prefer one grocer, the less they prefer the other. Thus, preference at one chain is a strong negative predictor of preference at the other. Stated differently, preference is specific to the chain, not shared within the format. This stands in contrast to the three mass merchandisers, for whom preference at one chain is a significant positive predictor of preference at the other two ($0.137 \leq \rho \leq 0.309$). Thus, a preference to patronize mass merchandisers is shared among store chains within the format.

Next, we consider preference correlations for the conditional spending decision in the lower panel of table 8. These correlations capture relationships in preference for spending at the six retailers. As with the patronage decision, we find a significant negative relationship in preferences for the two grocery chains ($\rho = -0.363$). A preference for spending at one grocery chain is therefore a negative predictor of preference for spending at the other. This finding, together with the large negative correlation in patronage preferences, suggests an overall substitution relationship between the two grocery retailers. The remaining conditional spending correlations appear to reflect a household-level preference for promotional levels across stores. We find that spending preferences are positively correlated ($\rho = 0.217$) at the two most promotional retailers, the drug store chain and grocery 1 (mean promotional discounts 24.3% and 21.5%, respectively). Spending at the least promotional retailer, mass merchandiser

3 (mean promotional discount 14.1%), is largely uncorrelated with spending at other chains, though it is somewhat negatively correlated with grocery 2 ($\rho = -0.127$) and mass merchandiser 1 ($\rho = -0.185$). Between these promotional extremes, spending at grocery 2 and mass merchandisers 1 and 2 are all positively correlated and highly significant ($0.246 \leq \rho \leq 0.518$). Thus, spending preferences seem to reflect “promotional tiers,” in this case deep, moderate, and shallow discounters.

In summary, patronage preferences are largely correlated within formats, with negative intraformat correlations between grocery stores and positive intraformat correlations among mass merchandisers. By contrast, spending correlations reflect a different retail segmentation, perhaps based on promotional discounting policies. We also note the significant positive correlations between the drug store chain and mass merchandiser 1 for both patronage ($\rho = 0.210$) and spending ($\rho = 0.563$) preferences. We conjecture that this affinity may result from the fact that both chains have a large number of stores in urban areas.

Next, we examine preference correlations at the format level. To do so, household-level intercepts are pooled across chains for each format as follows (patronage models are shown for exposition):

$$M\mathbf{t}_i \sim \text{MVN}(0, \text{MPM}'), \quad (13)$$

where $\mathbf{0} = (0 \ 0 \ \dots \ 0)'$, $\mathbf{t}_i \sim \text{MVN}(0, \text{P})$,¹² and

$$M = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 \end{bmatrix}. \quad (14)$$

Premultiplying by the matrix, M , sums the intercepts in each of the three formats. Cross-format correlations are computed from the adjusted covariance matrix, MPM' . Because α 's (from the conditional spending specifications) represent expectations of the logarithm of spending, we must exponentiate α 's before pooling and then take the logarithm of the sum.

Table 9 shows preference correlations between formats. Because our analysis of store-level patronage preferences found correlations primarily within formats, it is not surprising that there are no significant correlations in patronage preferences between formats. Cross-format correlations in spending preferences are all positive, and two are significant. In particular, we find a positive correlation ($\rho = 0.363$) in spending between grocers and mass merchandisers. While we certainly cannot say that these formats do not compete, we can predict that households that prefer to spend more at grocery stores will also prefer to spend more at mass merchandisers. It is worth noting that all cross-format preference correlations, for both patronage and spending, are positive. This suggests a general household-level preference for shopping that goes across formats.

12. The α_{hi} 's and t_{hi} 's estimated in the hierarchy have a zero-mean because $\bar{\alpha}$, and \bar{t} , are estimated during the first stage.

TABLE 9 Preference Correlations across Formats

	Drug Store	Grocery Store	Mass Merchandiser
Patronage—where to shop:			
Drug store	1	.072 (.077)	.073 (.044)
Grocery store		1	.090 (.078)
Mass merchandiser			1
Expenditure—how much to spend:			
Drug store	1	.161** (.065)	.123 (.094)
Grocery store		1	.363*** (.064)
Mass merchandiser			1

NOTE.—Standard errors are shown in parentheses below the estimates.

* Statistically significant at .05.

** Statistically significant at .01.

*** Statistically significant at .001.

V. Discussion and Managerial Implications

We began this article by noting that grocery retailers view mass merchandisers as a competitive threat. Our analysis of unexplained expenditures (see table 9) does not show a direct substitution relationship, even though the products sold at mass merchandisers overlap with traditional grocers. Households that prefer to spend more at grocery stores also prefer to spend more at mass merchandisers. Moreover, the negative preference correlations for patronage and conditional spending between grocery retailers suggest that substitution within the grocery format is much stronger than substitution between grocery and nongrocery formats.

To illustrate the implications of our findings regarding response to the marketing mix, we consider two possible strategic objectives for a grocery retailer: (1) to increase its customer base by attracting more shoppers and (2) to increase spending per customer in its stores. Achieving either of these objectives will result in higher revenues and, depending on costs, higher profits for the chain. We conduct a sensitivity analysis to show how these objectives might be achieved by changing marketing policies. Specifically, we use our model to predict how the retailer's customer base and revenues would respond to more aggressive levels of promotion, assortment, and market penetration (i.e., more stores). For example, how would grocery revenues respond if grocers sold 2% more of their items on promotion? What if the retailer offered assortments in every category that were 2% deeper? What if the retailer offered 5% more stores in the market area, and shoppers' travel times were correspondingly shorter? We note, however, that the range of the data for these variables is limited, making predictions beyond a narrow range problematic. Table 10 shows the percentage of households shopping (i.e., customer base) and share of revenues among the stores in our data set for grocers 1 and 2, both with current and hypothetical levels of the marketing variables. Note that deltas in

TABLE 10 Sensitivity Analysis—Marketing Variables

	Actual Baseline (%)	Δ Expected (%) If Promotion Increased by 2%	Δ Expected (%) If Assortment Increased by 2%	Δ Expected (%) If Penetration Increased by 5%
Grocery 1:				
Households shopping (%)	58.20	-.59	3.25	5.68
Share of revenues	23.64	-.95	8.08	2.45
Grocery 2:				
Households shopping (%)	83.59	.36	2.17	.52
Share of revenues	45.11	1.39	5.19	.77

table 10 are proportional changes in baseline levels of shoppers and market share.

If either grocery chain were to offer deeper assortments or locate more stores in the market area, their customer base and share of revenues would increase. Shoppers are highly sensitive to assortments at both grocery chains. If grocery 1 were to increase its assortments by 2% across all categories, its customer base would increase to 60.09% of all households (a proportional increase of 3.25%), and its market share would increase by 1.91 share points (a proportional increase of 8.08%). Grocery 2 would also benefit from a similar increase in assortment, though somewhat less so. A 2% assortment increase would raise its customer base to 84.51% of households (a proportional increase of 2.17%) and augment its market share by 2.34 share points (a proportional improvement of 5.19%). Increasing market penetration by 5% for both stores would result in more modest improvements in market share but differential effects on customer base. If grocery 1 were to increase its market penetration by 5% (four stores), its customer base would increase dramatically to 61.51% of all households (a proportional increase of 5.68%), while its market share would increase by only 0.58 share points (a proportional increase of 2.45%). Thus, the average customer would spend less if grocery 1 were to increase its penetration. Grocery 2 would benefit far less by increasing its penetration by 5% (nine stores), perhaps because its current high penetration results in a ceiling effect. Greater market penetration for grocery 2 would result in a customer base increase to 84.03% (a proportional improvement of only 0.52%) and a market share gain of only 0.35 share points (a proportional increase of only 0.77%).

Increasing promotions would have a differential effect on the two grocery retailers. Were they unilaterally to increase promotions, the smaller, more promotional grocery 1 would suffer, while the larger, less promotional grocery 2 would benefit. An increase in promotional intensity at grocery 1 such that 2% more purchases are made on discounted items would result in the customer base shrinking slightly to 57.86% of households (a proportional decrease of 0.59%), with a loss of 0.23 share points (a proportional decrease of 0.95%). A similar increase in promotional intensity at grocery 2 would result in augmenting the customer base to 83.90% of all households (a proportional gain

of 0.36%) and raising market share by 0.63 points (a proportional gain of 1.39%).

Should these grocers focus their efforts on building their customer base or increasing customer spending? Grocery 1 has a relatively smaller number of stores (84 in our market area) and a relatively smaller customer base (58.20%). We find that the benefits of increasing penetration and, to a lesser extent, increasing assortment depend on attracting new customers. Thus, expanding its customer base is an appropriate objective for grocery 1. This is not the case for grocery 2. This retailer stands to gain few customers from increasing its market penetration (through adding stores to its current total of 171), or by offering more promotions. Grocery 2 could, however, augment its customer base by extending category assortments. Both retailers would realize substantial increases in spending per customer by offering deeper assortments, with market shares increasing at a proportionally higher rate than customer bases. Grocery 2 could also raise spending per customer by increasing promotional intensity. In sum, the more appropriate strategic objective for grocery 2 is to focus on sales per customer.

How cost effective is it to address these objectives by offering deep assortments? If we assume that each store carries about 25,000 stock keeping units, then the incremental customers and revenues are available at a cost of inventorying and merchandising only 500 more products ($25,000 \times 2\%$). This is a relatively low cost, particularly when compared with adding stores, and with the attendant real estate, inventory, labor, and overhead costs. Note that we have not modeled out-of-stocks, which might increase were assortments extended in each category. Out-of-stocks would certainly have a negative impact on patronage and spending.

VI. Summary and Future Research Directions

This research represents the first study of household-level shopping behavior across retail formats. The hierarchical multivariate type-2 tobit model introduced in this article provides a flexible framework with which to analyze shopper's decisions about where to shop and how much to spend. We summarize our empirical findings below and highlight hypotheses for future research:

1. Much of the variability in expenditures across formats can be explained by retail format alone. In fact, 31.1% of the variation in household-level monthly expenditures across formats can be explained with a model specifying only retailer intercepts (in a system of independent type-2 tobits with fixed effects).¹³ This is likely because of the very different marketing policies that each of these formats follows (see table 1). An interesting direction for future

13. This figure reported is a pseudo R^2 computed from the residuals of the specified system of tobit models.

cross-format research would be to investigate whether consumers' shopping processes have higher-order shopping strategies (i.e., they anticipate long-term needs when making decisions about where to shop).

2. Among marketing variables, store patronage and spending are highly responsive to differences in retailers' promotional intensity, both over time and across shoppers' market baskets. Future research that addresses why we observe differential response to promotions across formats, and how retailers should therefore change promotional strategies, would be useful. Household patronage and spending are also sensitive to differences in retailer assortments across market baskets, particularly at grocery chains. This is somewhat surprising given the grocery industry's focus on finding ways to reduce category-level assortments. Future research on this topic must determine the effect of assortment on patronage or store choice, rather than category sales alone.

3. Consumers' store-level shopping decisions are insensitive to monthly variation in the price of a market basket. We have found that relative basket prices are extremely stable over time (coefficient of variation = 0.022), and consumers are either unable to discern these small changes in basket prices (Alba et al. 1994) or are not troubled to act on them (Kalyanaram and Little 1994). Estimated revenue elasticities for all formats are not significantly different from zero, indicating that quantity price elasticities are not significantly different from unity. This suggests the hypothesis that consumer spending is insensitive to observed variation in market-basket prices for packaged goods retailers. More general future research, including research on other markets or nonpackaged goods retailers, would offer useful tests of this hypothesis and its limitations.

4. Of the formats considered, mass merchandisers are least sensitive to shoppers' travel time. Future research might address optimal store placement (both where and how many), including differences between formats.

5. The two previous findings can be illuminated by relating mean levels of marketing variables and consumer response to cross-sectional (and in some cases temporal) variation in those levels. Revenue elasticities of promotion ($\eta_{\text{drug}} = -1.323$, $\eta_{\text{gro.}} = 0.633$, and $\eta_{\text{mass}} = 0.729$) are inversely related to mean levels of both advertised discount ($\text{discount}_{\text{drug}} = 24.3\%$, $\text{discount}_{\text{gro.}} = 20.1\%$, and $\text{discount}_{\text{mass}} = 16.7\%$) and percentage of sales on promotion ($\% \text{sales}_{\text{drug}} = 29.3$, $\% \text{sales}_{\text{gro.}} = 18.2$, and $\% \text{sales}_{\text{mass}} = 13.9$). Thus, less promotional formats could realize greater revenues by increasing promotional intensity, while more promotional formats would benefit by reducing promotions. By contrast, revenue elasticities of assortment ($\eta_{\text{drug}} = -0.109$, $\eta_{\text{gro.}} = 4.972$, and $\eta_{\text{mass}} = 0.897$) are positively related to mean assortment indices ($\text{assortidx}_{\text{drug}} = 0.408$, $\text{assortidx}_{\text{gro.}} = 1.868$, and $\text{assortidx}_{\text{mass}} = 0.587$). This indicates that formats with greater assortments could benefit by increasing their offerings, while the lowest assortment format could not. In other words, while retailers across formats are responding to consumers' sensitivity to assortment, grocers in particular would benefit from offering even deeper assortments. We also find that

revenue elasticities of travel ($\eta_{\text{drug}} = -0.243$, $\eta_{\text{gro.}} = -0.308$, and $\eta_{\text{mass}} = -0.106$) are somewhat related to mean levels of travel time ($\text{traveltime}_{\text{drug}} = 9.6$, $\text{traveltime}_{\text{gro.}} = 11.0$, and $\text{traveltime}_{\text{mass}} = 15.6$), which result from retailers' market penetration strategies. We find that the lower penetration of mass merchandisers is consistent with shoppers' insensitivity to travel for this format. We believe that these travel time elasticities are biased downward because of measurement error resulting from trip chaining. Future research concerning how retailers could more effectively respond to differential sensitivity to the market mix across formats would be important.

6. Households that have higher intrinsic preferences for spending at grocery stores also prefer to spend more at other formats, particularly mass merchandisers. Within grocery stores, spending preferences are negatively related. It also appears that spending preferences exist within promotional "tiers," that is, spending preferences at the most promotional stores (drug chain and grocery 1) are positively related, while preferences at moderately promotional stores (grocery 2, mass merchandisers 1 and 2) are also positively related. No strong relationships exist for patronage preferences across formats, though relationships within formats are apparent. Between the two grocery stores we study, an intrinsic preference for patronizing one chain is a strong negative predictor of preference for the other. Among mass merchandisers, a preference to patronize one chain is a positive predictor of preference for the others. Taken together, these findings indicate that competition between formats is fundamentally different than competition within formats and suggest that, across formats, stores are not close substitutes. Studies that include multiple geographic area and larger panels are needed to verify our initial findings about competition across retail formats. While recent work has considered why consumers choose different stores on different trips, little has been done to determine the substitutability or complementarity of stores of different formats. In addition, we offer the hypothesis for future study that, across formats, consumers prefer to shop at stores with similar promotional policies.

In conclusion, we hope that our results will foster more research in the area of cross-format shopping. We must point out the limitations of our study, however. Our data come from only one metropolitan market over a 2-year span. Other markets may exhibit different characteristics, and the time span of our data may not be long enough to capture long-term trends. In addition, we have focused on the household's aggregate purchases across all categories on a monthly basis. We believe that separating products into those categories that are carried in common across the stores (e.g., dry packaged groceries) and those that are not (e.g., produce, meat, and bakery items) would be an important advance over our current research. Unfortunately, this increases the dimension of the problem substantially. Further, it is likely that some households are using higher-order shopping strategies, that is, visiting multiple stores on a single shopping visit or dynamically determining "stock-up" and "fill-in" trips. Taking a more holistic approach to consumption, purchases, time

allocation, and the household production function would also greatly advance our understanding of cross-format shopping. It is our hope that the results from this study will aid future research on these topics.

Appendix A

Aggregation Effects on Price Variation

In order to assess the impact of aggregation on price variation, we evaluate prices at different levels of aggregation. We define aggregation along two dimensions—products and time. Aggregation over products is intended to reflect shopping decisions at different levels. For brand choice and quantity decisions, UPC or item-level prices are germane, so this provides our baseline. For category incidence, prices are typically evaluated at the category level (e.g., Dillon and Gupta [1996], another study in which item prices are weighted components of category attractiveness). For store-level shopping decisions, the relevant price is for a basket of products, including many categories (e.g., Bell, Ho, and Tang 1998; Bell, Bucklin, and Sismiero 2000). Temporal aggregation levels are weekly and monthly. Retailer prices change weekly, so this is the usual baseline level (although in fact a very small percentage of prices are changed each week). Our PRICE variable, used to predict store-level patronage and quantity, is aggregated at the monthly level.

To examine price variation at different levels of aggregation, we compute “true” prices from merchandise files for all items that are sold at the two grocery stores and one drug store chain in our data set, across nine categories (2,262 total UPCs).¹⁴ The coefficient of variation (standard deviation divided by mean) is computed to measure price variation. For UPC and category cross-sectional levels, the coefficients of variation are computed at that level, then averaged over the 2,262 products or nine categories, respectively. Note that our small market basket captures prices in only a few of the 261 packaged goods categories and so does not reflect the true breadth of the average market basket. As such, it represents a conservative test of the effect of product aggregation at the basket level. The three store chains provide repeated measures of price variation. Results are reported in table A1.

We observe that, in general, there is less variation in prices computed on a monthly basis than on a weekly one. However, as cross-sectional aggregation increases, the effect of temporal aggregation is attenuated. In fact, more than 95% of the price variation for the small market basket is preserved when aggregating prices from weekly to monthly (across the three retailers). By contrast, price variation decreases quite rapidly when aggregating from the “category” to the “small market basket” level (less than 42% of price variation is retained). An ANOVA of the data in the table shows very clearly that cross-sectional, or product, aggregation has an order-of-magnitude larger effect on the coefficient of variation (mean square = 0.0027) compared with temporal aggregation (mean square = 0.0002). These two factors together explain the coefficient of variation quite well, with $R^2 = 0.947$.

An important benefit of computing basket prices and promotional intensity across a great many categories is that we need make no assumptions about the representa-

14. Among the stores captured in the analysis, merchandise files are available for only these three retailers.

TABLE A1 Cross-Sectional Aggregation

	UPC			Category			Small Market Basket ^a		
	Drug Store	Grocery 1	Grocery 2	Drug Store	Grocery 1	Grocery 2	Drug Store	Grocery 1	Grocery 2
Temporal (weekly)	.0642	.0718	.0597	.0747	.0713	.0747	.0308	.0267	.0342
Aggregation (monthly)	.0487	.0532	.0502	.0706	.0659	.0706	.0297	.0248	.0331

^a UPC = uniform product codes. The small market basket comprises products in nine categories: beer and ale, chocolate candy, salty snacks, internal analgesics, sanitary napkins, cigarettes, diapers, dog food, and household cleaners.

tiveness of a small number of categories for the shopper's entire basket. Thus, our individually weighted measure incorporating all packaged goods categories should more accurately reflect the true basket prices and promotions faced by the shopper. However, we do implicitly assume that the relative price, promotion, and assortment levels of packaged goods reflect nonpackaged goods as well (e.g., perishables such as produce and meat, clothing, and durable items). Prices of such items are not available in syndicated panel data.

Appendix B

Using Residual Correlations for Prediction

The residual covariances have predictive value. Note that the conditional distribution for patronage of chain i is:

$$\Pr(z_{hit} = 1 | z_{hkt}^*) = \Phi\left(\nu_{hi} + \mathbf{x}_{hit}\theta'_i + d_{hi}\chi_i + \mathbf{s}'_h\kappa_i + \frac{\lambda_{ik}}{\lambda_{kk}}[z_{hkt}^* - E(z_{hkt}^*)]\right), \quad (\text{B1})$$

where $i \neq k$ and $\lambda_{ik} = (\Lambda)_{ik}$. Suppose we know that, if a household frequents mass merchandiser 3 more than expected, the probability of frequenting grocery 2 is also higher than expected. More precisely, suppose that the unconditional probability that a household visits grocery 2 is 50%. If we know that this household has a higher probability of visiting mass merchandiser 3 than expected, for example, let $[z_{bk}^* - E(z_{bk}^*)]/(\sqrt{\lambda_{kk}} = 2)$, then the conditional probability of this household visiting grocery 2 is 65%.

Alternatively, if the probability of shopping at all stores except the one of interest is known, we can compute the reduction in standard deviation of z_{hi}^* to assess how this additional information can improve the prediction of whether a household will come to the chain of interest. For example, if we are interested in predicting patronage at mass merchandiser 3, and we know patronage at all other stores, the standard deviation of z_{hi}^* will be reduced by 31%. The range of reductions goes between 17% and 31% for the six chains in this study. Clearly, deviations from expected patronage at other stores can be very valuable to a retailer in predicting whether an individual will shop at their store.

References

- Ainslie, Andrew, and Peter E. Rossi. 1998. Similarities in choice behavior across product categories. *Marketing Science* 17 (Spring): 91–106.
- Alba, Joseph W., Carl F. Mela, Terence A. Shimp, and Joel E. Urbany. 1994. The influence of prior beliefs, frequency cues, and magnitude cues on consumer price judgments. *Journal of Consumer Research* 26 (September): 99–114.
- Amemiya, Takeshi. 1985. *Advanced econometrics*. Cambridge, MA: Harvard University Press.
- Arentze, Theo A., Aloys W. J. Borgers, and Harry J. P. Timmermans. 1993. A model of multi-purpose shopping trip behavior. *Papers in Regional Science* 72, no. 3:239–56.
- Arnold, Stephen J., S. Ma, and Douglas J. Tigert. 1978. A comparative analysis of determinant attributes in retail store selection. In *Advances in consumer research*, vol. 5, ed. H. K. Hunt. Ann Arbor, MI: Association for Consumer Research.
- Arnold, Stephen J., Tae H. Oum, and Douglas J. Tigert. 1983. Determinant attributes in retail patronage: Seasonal, temporal, regional and international comparisons. *Journal of Marketing Research* 20 (May): 149–57.

- Arnold, Stephen J., and Douglas J. Tigert. 1982. Comparative analysis of determinants of patronage. In *Retail patronage theory: 1981 workshop proceedings*, ed. Robert F. Lusch and William R. Darden. Norman: University of Oklahoma Printing Services.
- Barnard, Peter O., and David A. Hensher. 1992. The spatial distribution of retail expenditures. *Journal of Transport Economics and Policy* 26 (September): 299–312.
- Bell, David R., Randolph E. Bucklin, and Catarina Sismiero. 2000. Consumer shopping behaviors and in-store expenditure decisions. Working paper, University of Pennsylvania, Wharton School, Philadelphia.
- Bell, David R., Teck-Hua Ho, and Christopher S. Tang. 1998. Determining where to shop: Fixed and variable costs of shopping. *Journal of Marketing Research* 35 (August): 352–69.
- Bell, David R., and James M. Lattin. 1998. Shopping behavior and consumer preference for store price format: Why “large basket” shoppers prefer EDLP. *Marketing Science* 17 (Winter): 66–88.
- Blattberg, Robert C., Richard Briesch, and Edward J. Fox. 1995. How promotions work. *Marketing Science* 14 (Summer): G122–G132.
- Blattberg, Robert C., Thomas Buesing, Peter Peacock, and Subrata K. Sen. 1978. Identifying the deal prone segment. *Journal of Marketing Research* 15 (August): 369–77.
- Blundell, Richard W., and Richard J. Smith. 1994. Coherency and estimation of simultaneous models with censored or qualitative dependent variables. *Journal of Econometrics* 64:355–73.
- Boatwright, Peter, and Joseph C. Nunes. 2001. Reducing assortment: An attribute-based approach. *Journal of Marketing* 65 (July): 50–63.
- Brown, Stephen. 1989. Retail location theory: The legacy of Harold Hotelling. *Journal of Retailing* 65 (Winter): 450–70.
- Bucklin, Randolph E., and Sunil Gupta. 1992. Brand choice, purchase incidence and segmentation: An integrated approach. *Journal of Marketing Research* 29:201–15.
- Casella, George, and Edward I. George. 1992. Explaining the Gibbs sampler. *American Statistician* 46 (August): 167–74.
- Chib, Siddhartha. 1993. Bayes inference in the tobit censored regression model. *Journal of Econometrics* 51:79–99.
- Chintagunta, Pradeep K., and Sunil Gupta. 1994. On using demographic variables to determine segment membership in logit mixture models. *Journal of Marketing Research* 31 (February): 128–39.
- Cornick, Jorge, Thomas L. Cox, and Brian W. Gould. 1994. Fluid milk purchases: A multivariate tobit analysis. *American Journal of Agricultural Economics* 76 (February): 74–82.
- Dellaert, Benedict G. C., Theo A. Arentze, Michel Bierlaire, Aloys W. J. Borgers, and Harry J. P. Timmermans. 1998. Investigating consumers’ tendency to combine multiple shopping purposes and destinations. *Journal of Marketing Research* 35 (May): 177–88.
- Dillon, William R., and Sunil Gupta. 1996. A segment-level model of category volume and brand choice. *Marketing Science* 15 (Winter): 38–59.
- Discount Store News*. 1999. The DSN annual industry report: Mass market retail breaks \$400 billion barrier. *Discount Store News* 38, no. 13:51.
- Food and Beverage Marketing*. 1994. Kroger’s total store challenge. *Food and Beverage Marketing* (May), pp. 18–34.
- Fox, Edward J., Richard Metters, and John Semple. 2003. Every house a warehouse: An inventory-theoretic model of retail shopping behavior. Working paper, Southern Methodist University, Marketing Department, Dallas.
- Gelfand, Alan E., and Adrian F. M. Smith. 1990. Sampling-based approaches to calculating marginal densities. *Journal of the American Statistical Association* 85:398–409.
- Ho, Teck-Hua, Christopher S. Tang, and David R. Bell. 1998. Rational shopping behavior and the option value of variable pricing. *Management Science* 44 (December): S145–S160.
- Hoch, Stephen J., Xavier Drèze, and Mary E. Purk. 1994. EDLP, hi-lo, and margin arithmetic. *Journal of Marketing* 58:16–27.
- Hoch, Stephen J., Byung D. Kim, Alan L. Montgomery, and Peter E. Rossi. 1995. Determinants of store-level price elasticity. *Journal of Marketing Research* 32 (February): 17–29.
- Hotelling, Harold. 1929. Stability in competition. *Economic Journal* 39 (March): 41–57.
- Hubbard, R. 1978. A review of selected factors conditioning consumer travel behavior. *Journal of Consumer Research* 5 (June): 1–21.
- Huff, David. 1964. Redefining and estimating a trading area. *Journal of Marketing* 28:34–38.
- Information Resources, and Willard Bishop Consulting. 1993. *Variety or duplication: A process to know where you stand*. Washington, DC: Food Marketing Institute.

- Jain, Dipak C., Naufel J. Vilcassim, and Pradeep K. Chintagunta. 1994. A random coefficients logit brand-choice model applied to panel data. *Journal of Business and Economic Statistics* 12 (July): 317–28.
- Kalyanaram, Gurumurthy, and John D. C. Little. 1994. An empirical analysis of latitude of price acceptance in consumer packaged goods. *Journal of Consumer Research* 26 (December): 408–18.
- Kumar, V., and Robert V. Leone. 1988. Measuring the effect of retail store promotions on brand and store substitution. *Journal of Marketing Research* 25 (May): 178–85.
- Levy, Michael, and Barton A. Weitz. 2003. *Retail management*. 5th ed. Chicago: McGraw-Hill/Irwin.
- Man, Kasing S. 2004. Linear prediction of temporal aggregates under model misspecification. *International Journal of Forecasting*. In press.
- Merrefield, Dan. 1995. Building profit in supermarkets by reducing stockkeeping units. *Supermarket News* 25:2.
- Messinger, Paul R., and Chakravarti Narasimhan. 1997. A model of retail formats based on consumers' economizing on shopping time. *Marketing Science* 16 (Winter): 1–23.
- Progressive Grocer Report of the Grocery Industry*. 1999. Markets in motion. *Progressive Grocer Report of the Grocery Industry* (April), p. 9.
- Reilly, William J. 1931. *The law of retail gravitation*. New York: Knickerbocker.
- Rossi, Peter E., Robert McCulloch, and Greg Allenby. 1996. On the value of household purchase history information in target marketing. *Marketing Science* 15 (Winter): 321–40.
- Thill, Jean-Claude, and Isabelle Thomas. 1987. Toward conceptualizing trip chaining behavior: A review. *Geographical Analysis* 1:1–17.
- Tobin, James. 1958. Estimation of relationships for limited dependent variables. *Econometrica* 26:24–36.
- Walters, Rockney G. 1991. Assessing the impact of retail price promotions on product substitution, complementarity, purchase and interstore sales displacement. *Journal of Marketing* 55 (April): 17–28.
- Walters, Rockney G., and Scott B. MacKenzie. 1988. A structural equations analysis of the impact of price promotions on store performance. *Journal of Marketing Research* 25 (February): 51–63.
- Walters, Rockney G., and Heikki J. Rinne. 1986. An empirical investigation into the impact of price promotions on retail store performance. *Journal of Retailing* 62, no. 3:237–66.
- Zellner, Arnold. 1962. An efficient method of estimating seemingly unrelated regressions and tests for aggregation bias. *Journal of the American Statistical Association* 57:348–68.